

Summary

- We propose a new derivative-free algorithm based on Gaussian process optimisation and particle filters.
- Enables estimation of parameters in general state space models.
- Parameter estimates close to the true values are obtained using only a few samples from the log-likelihood.

Frequentistic parameter inference

We are interested in solving the **parameter inference** problem in **nonlinear state space models**

$$\begin{aligned} x_{t+1}|x_t &\sim f_\theta(x_{t+1}|x_t), \\ y_t|x_t &\sim g_\theta(y_t|x_t), \end{aligned}$$

given a set of observations $\mathcal{D}_T = \{y_t\}_{t=1}^T$ and where $\theta \in \Theta \subseteq \mathbb{R}^d$ denotes static parameters. The **maximum likelihood estimate** is given by

$$\hat{\theta}_{\text{ML}} = \underset{\theta}{\operatorname{argmax}} \log p(\mathcal{D}_T|\theta),$$

where $\log p(\mathcal{D}_T|\theta)$ denotes the (often) intractable log-likelihood function.

Particle GP optimisation

In Bayesian optimisation, we iteratively optimise a **surrogate function** $f(\theta)$ modelled as a Gaussian process by a three step procedure.

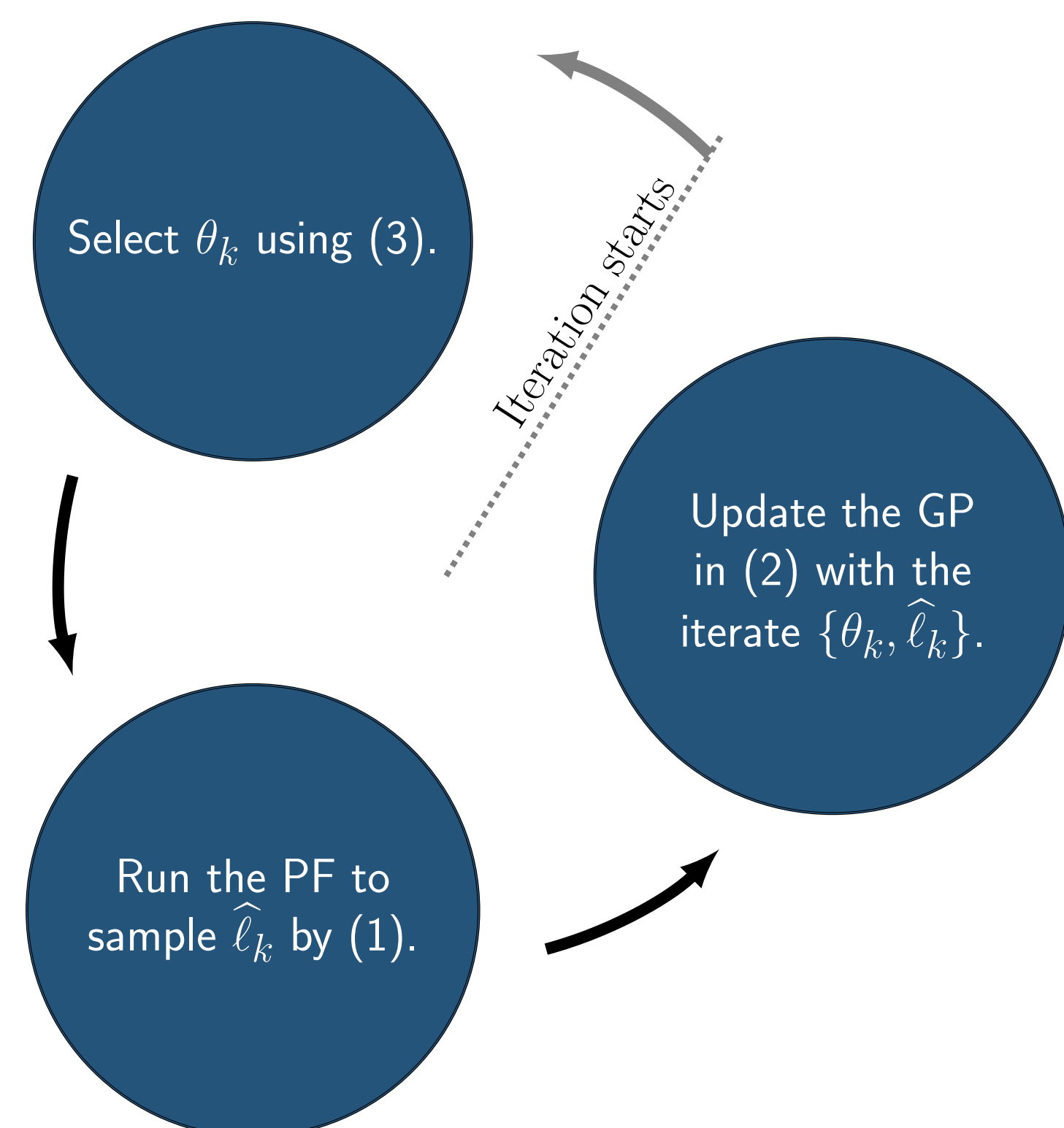


Figure: An iteration of the particle GP optimisation algorithm.

Main idea

Explore the likelihood landscape using a combination of particle filtering and Gaussian process models. New parameters are sampled according to the expected improvement of the model.

Estimating the likelihood

We run a **particle filter** (PF) targeting $p_{\theta_k}(x_t|\mathcal{D}_T)$, which returns the unnormalised particle system $\{x_t^{(i)}, w_t^{(i)}\}_{i=1}^N$. The log-likelihood can then be estimated using

$$\hat{\ell}_k = \log \hat{p}(\mathcal{D}_T|\theta_k) = \sum_{t=1}^T \log \left[\sum_{i=1}^N w_t^{(i)} \right] - T \log N. \quad (1)$$

The resulting pair $\{\theta_k, \hat{\ell}_k\}_{k=1}^m$ denotes the iterates of the algorithm.

Gaussian process model

The surrogate function in this optimisation is modelled as

$$f(\theta) \sim \mathcal{GP}(m(\theta), k(\theta, \theta')), \quad (2)$$

with a constant mean function, $m(\theta)$, and the Matérn covariance function, $k(\theta, \theta')$, with $\nu = 3/2$. The mean and variance of the model

$$\begin{aligned} \mu(\theta) &= \mathbb{E} \left[f(\theta) | \{\theta_k, \hat{\ell}_k\}_{k=1}^m \right], \\ \sigma^2(\theta) &= \mathbb{V} \left[f(\theta) | \{\theta_k, \hat{\ell}_k\}_{k=1}^m \right], \end{aligned}$$

are updated recursively using standard results.

Acquisition rule

The next point in which to sample $p(\mathcal{D}_T|\theta)$ is determined by the maximising argument of **the expected improvement** defined as

$$\begin{aligned} \mathbb{E}\mathbb{I}(\theta) &= \left[\mu(\theta) - \max_{\theta} \mu(\theta) - \xi \right] \Phi(Z) + \sigma(\theta) \phi(Z), \quad (3) \\ Z &= \frac{1}{\sigma(\theta)} \left[\mu(\theta) - \max_{\theta} \mu(\theta) - \xi \right], \end{aligned}$$

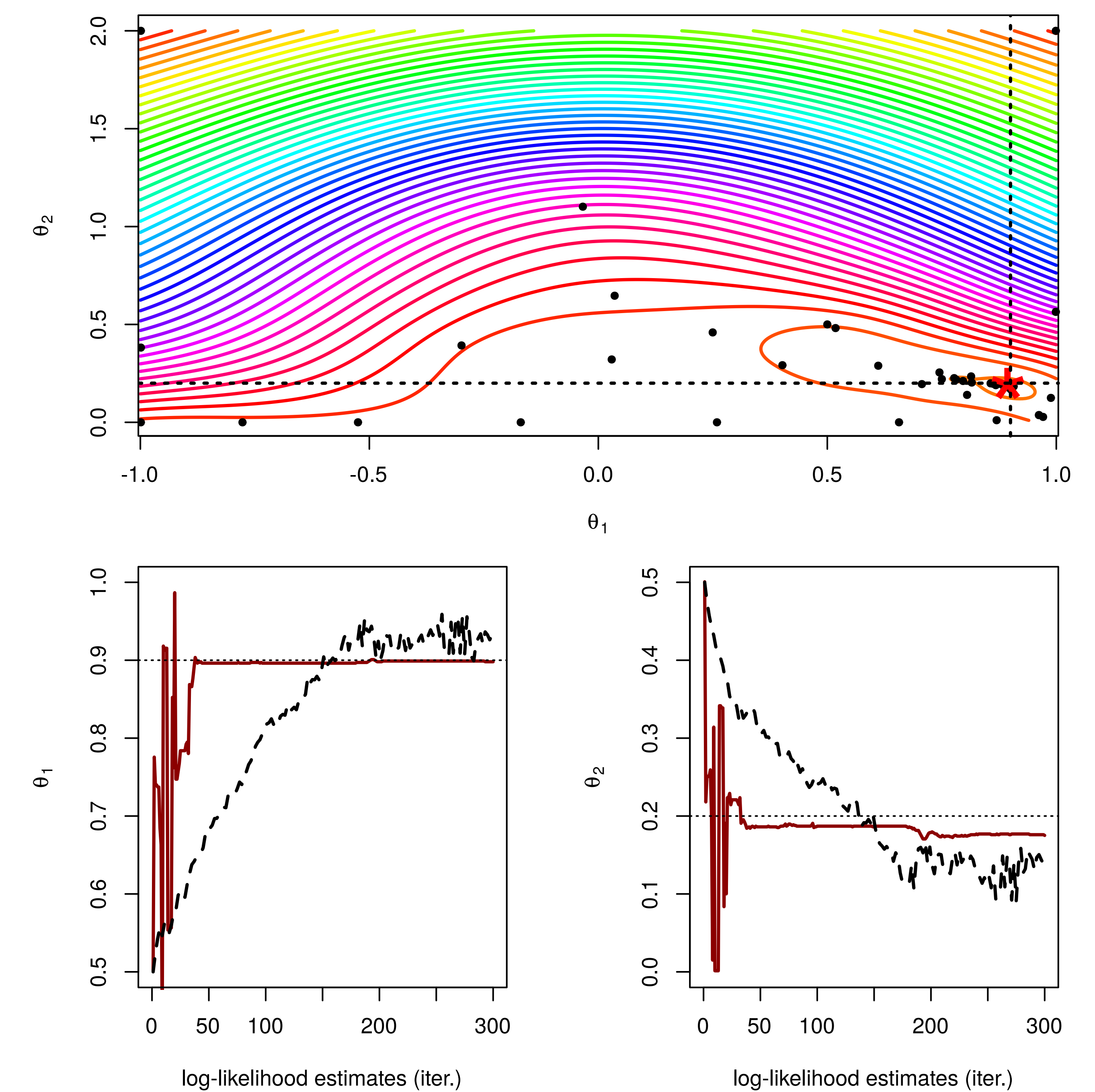
where ξ denotes a coefficient that balances exploration and exploitation. Here, Φ and ϕ denote the CDF and PDF of the Gaussian distribution.

Example: Stochastic volatility model

Consider the model

$$\begin{aligned} x_{t+1}|x_t &\sim \mathcal{N}(x_{t+1}; \theta_1 x_t, \theta_2^2), \\ y_t|x_t &\sim \mathcal{N}(y_t; 0, 0.7^2 \exp(x_t)), \end{aligned}$$

with true parameters $\theta^* = \{\theta_1^*, \theta_2^*\} = \{0.9, 0.2\}$. We use $T = 250$ time steps, $N = 1000$ particles and $M = 300$ iterations. The simultaneous perturbation stochastic approximation (SPSA) algorithm is used for a benchmark comparison.



Upper: the estimated GP posterior with the sampling points (solid dots), the true parameters (dashed lines) and the parameter estimate (red star). **Lower:** the parameter estimate at each iteration is presented for the GPO algorithm (red line) and the SPSA algorithm (black line).

More information and source code

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