

Constructing Metropolis-Hastings proposals using damped BFGS updates

The 18th IFAC Symposium on System Identification, Stockholm, Sweden.

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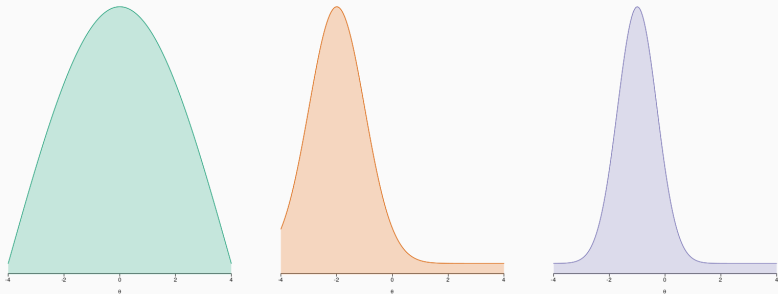
This is work in collaboration with

Dr. Adrian Wills (University of Newcastle, Australia)

Prof. Brett Ninness (University of Newcastle, Australia)



Bayesian inference in one slide



$$\pi(\theta) \triangleq p(\theta|y) = \frac{p(y|\theta)p(\theta)}{p(y)},$$

$$\pi[\varphi] \triangleq \mathbb{E}_{\pi}[\varphi(\theta)] = \int_{\Theta} \varphi(\theta)\pi(\theta) \, d\theta.$$

What are we going to do?

- Estimate posterior distributions using Markov chains.
- Improve the standard choice to handle high-dimensional problems.

Why are we doing this?

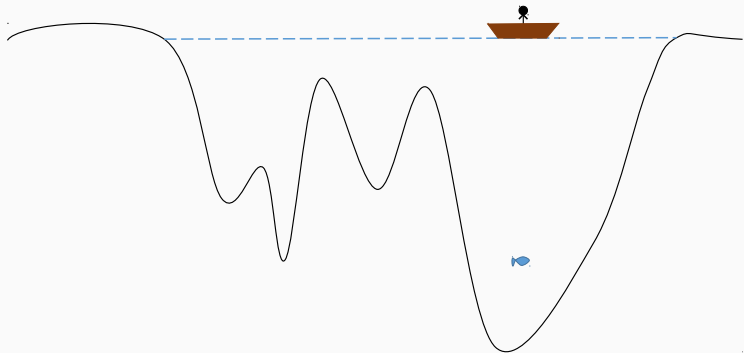
- Standard choice gives inefficient sampling: slow or not working at all.
- Bayesian methods give uncertainty and valid estimates for finite data.

How will we do this?

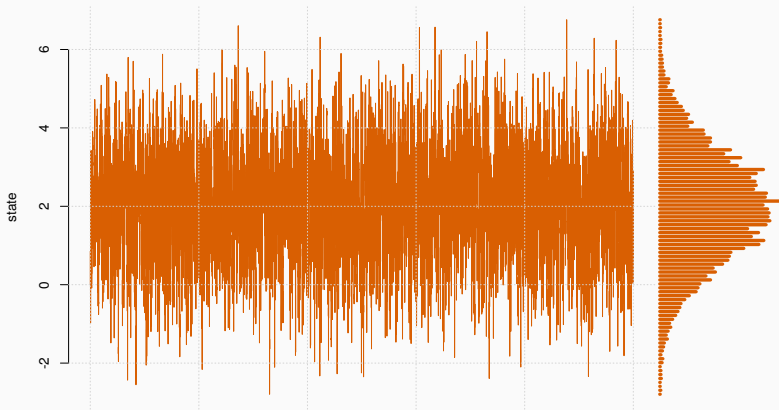
- Make use of gradient and Hessian information.
- Estimate the Hessian using quasi-Newton and least squares.

Exploring posteriors using Markov chains.

Exploring "the lake"



Markov chains: stationary distribution



Metropolis-Hastings: algorithm

Get samples from target $\pi(\theta) \propto p(y|\theta)p(\theta)$ by iterating over k :

(i) Propose **candidate parameter** θ' by

$$\theta' \sim q(\theta'|\theta_{k-1}).$$

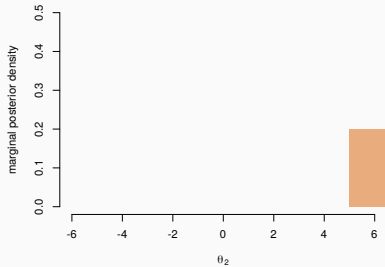
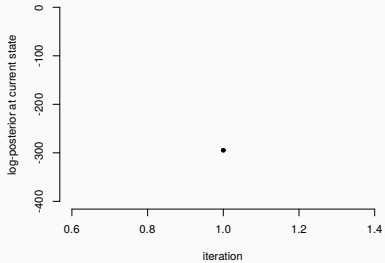
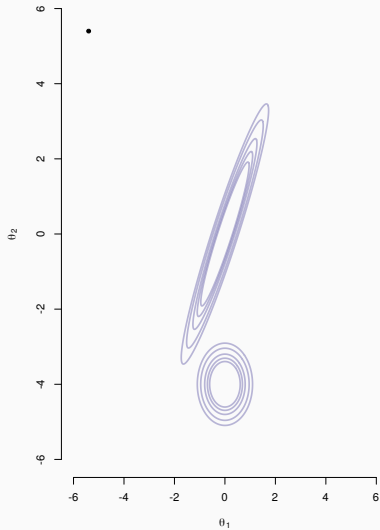
(ii) **Accept** θ' by setting $\theta_k \leftarrow \theta'$ with probability

$$\min \left\{ 1, \frac{\pi(\theta')}{\pi(\theta_{k-1})} \right\},$$

otherwise **reject** θ' by setting $\theta_k \leftarrow \theta_{k-1}$.

Output: Samples $\{\theta_k\}_{k=1}^K$ from $\pi(\theta)$.

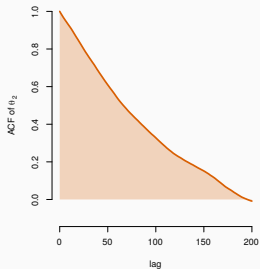
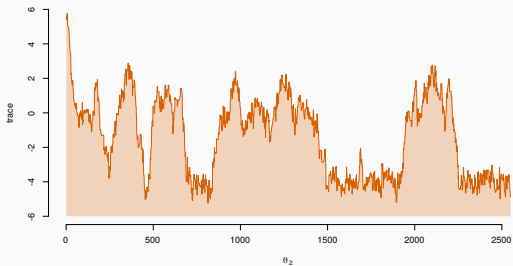
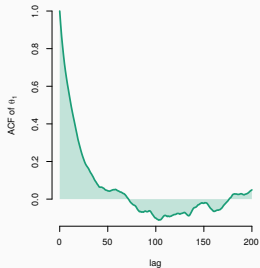
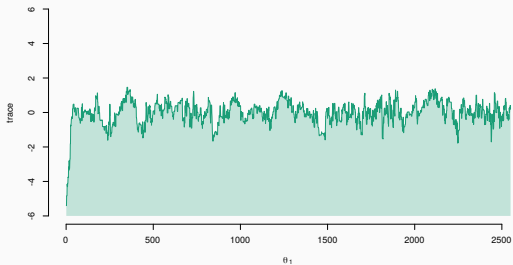
Metropolis-Hastings: toy example



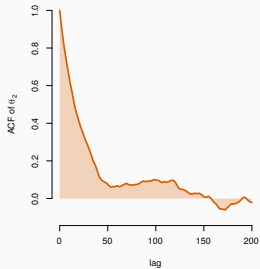
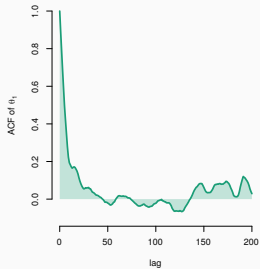
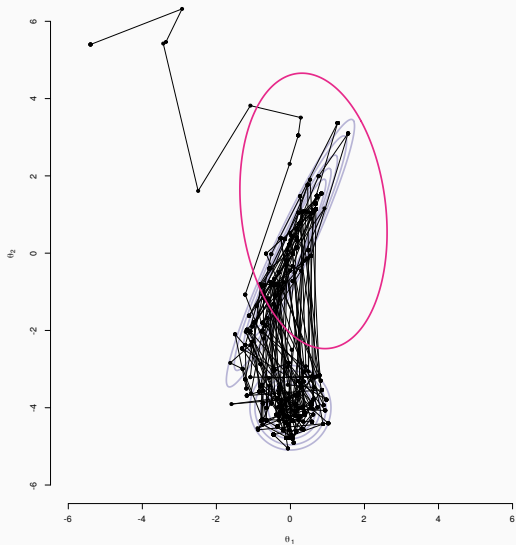
Metropolis-Hastings: toy example

Metropolis-Hastings: toy example

Metropolis-Hastings: proposal and mixing

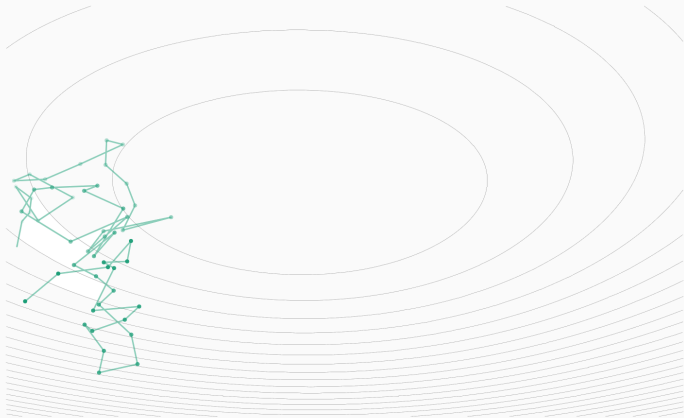


Metropolis-Hastings: proposal and mixing



Constructing efficient proposal distributions.

Gaussian random walk proposal



$$\theta' | \theta_{k-1} \sim \mathcal{N}(\theta' | \theta_{k-1}, \Sigma),$$

A second-order approximation

Second-order Taylor expansion of log-target

$$\log \pi(\theta + \Delta\theta) \approx \log \pi(\theta) + G(\theta)^\top \Delta\theta - \frac{1}{2} \Delta\theta^\top H(\theta) \Delta\theta,$$

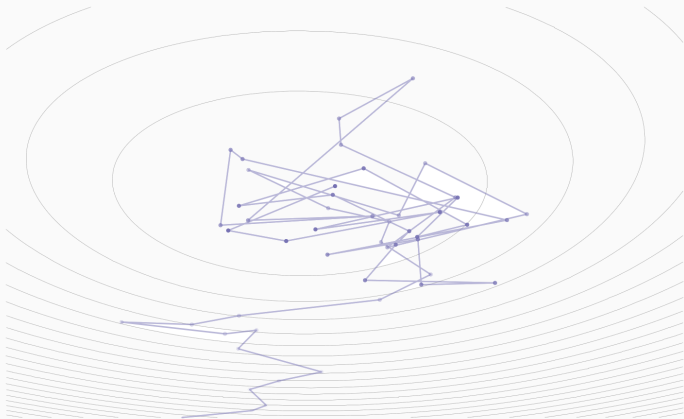
with the approximate gradient

$$\nabla \log \pi(\theta + \Delta\theta) \approx G(\theta) - H(\theta) \Delta\theta,$$

which by setting to zero gives the search direction

$$\Delta\theta = H(\theta)G(\theta).$$

Second-order Gaussian proposal



$$\theta' | \theta_{k-1} \sim \mathcal{N}\left(\theta' | \theta_{k-1} + \frac{1}{2} H(\theta_{k-1}) G(\theta_{k-1}), H(\theta_{k-1})\right).$$

Hessian estimation using quasi-Newton methods.

Hessian estimates using quasi-Newton

$$\text{BFGS} \quad \bar{H}_k = \bar{H}_{k-1} + \frac{s_k s_k^\top}{s_k^\top y_k} - \frac{\bar{H}_{k-1} y_k y_k^\top \bar{H}_{k-1}}{y_k^\top \bar{H}_{k-1} y_k},$$

$$\text{SR1} \quad \bar{H}_k = \bar{H}_{k-1} + \frac{(s_k - \bar{H}_{k-1} y_k)(s_k - \bar{H}_{k-1} y_k)^\top}{(s_k - \bar{H}_{k-1} y_k)^\top y_k},$$

$$s_k \triangleq \theta_k - \theta_{k-1}, \quad y_k \triangleq G(\theta_k) - G(\theta_{k-1}).$$

Hessian estimate using least squares

The quasi-Newton estimate $\bar{H}(\theta)$ is assumed to fulfill

$$G(\theta + \Delta\theta) = G(\theta) + \bar{H}(\theta)\Delta\theta,$$

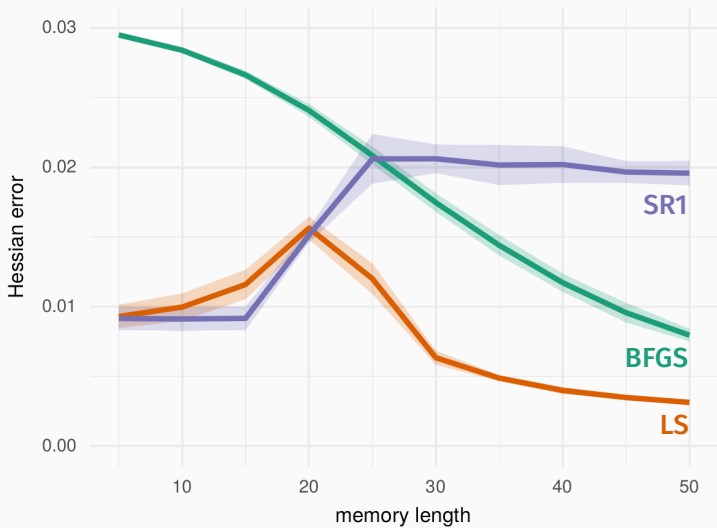
which is called the **secant condition**. Introduce

$$Y_k = [y_2, \dots, y_k], \quad S_k = [s_2, \dots, s_k],$$

and compute a least squares estimate of

$$Y_k = \bar{H}_k S_k.$$

Hessian approximation error



Numerical illustrations.

Detecting the Higgs boson

$N = 11,000,000$ events generated with binary observations (detection/background) given $p = 21$ kinematic properties.

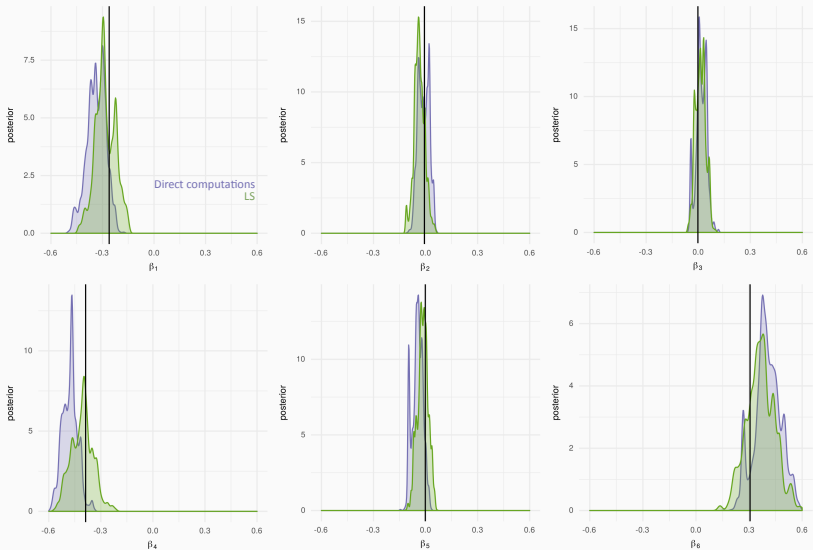
Modelling the data by a **logistic regression model**:

$$y_t \sim \text{Bernoulli}(p_t), \quad p_t = \text{logit} \left(\beta_0 + \sum_{i=1}^p \beta_i x_{it} \right).$$

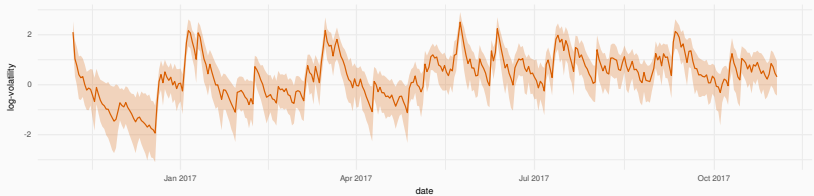
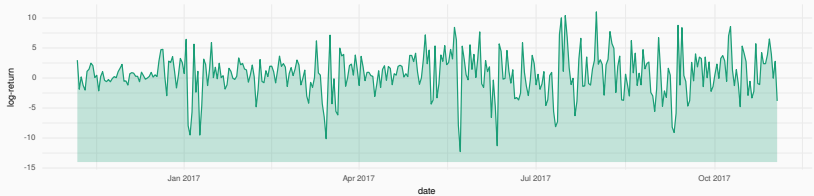
where the parameters are $\theta = \{\beta_0, \beta_1, \dots, \beta_{21}\}$.

Big data setting: **sub-sampling** methods are required.

Detecting the Higgs boson, cont.



Modelling Bitcoin prices



Modelling Bitcoin prices

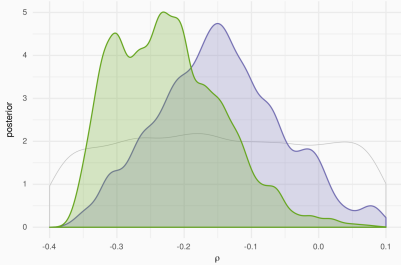
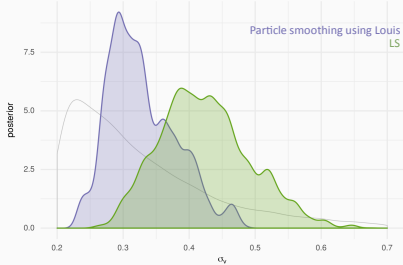
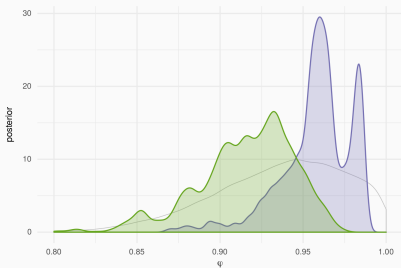
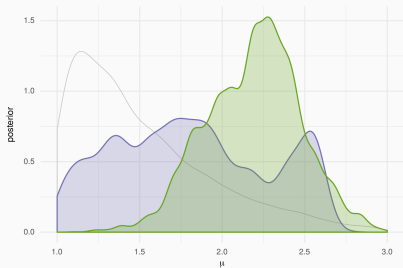
Modelling the data using a **stochastic volatility model**:

$$x_{t+1}|x_t \sim \text{N}\left(x_{t+1} \mid \mu + \phi(x_t - \mu) + \rho\sigma_v y_t \exp(-x_t), \sigma_v^2\right),$$
$$y_t|x_t \sim \text{N}\left(y_t \mid 0, \exp(x_t)\right)$$

where the parameters are $\theta = \{\mu, \phi, \sigma_v, \rho, x_{0:T}\}$.

Latent variables: **particle smoothing** is required.

Modelling Bitcoin prices, cont.



What did we do?

- Employed a second-order approximation as a proposal.
- Estimated the Hessian using quasi-Newton and least squares.
- Applied the approach to high-dimensional problems.

Why did we do this?

- Standard proposals does not scale well.
- Hessian estimates are difficult to compute directly.
- Bayesian methods work on finite data and gives uncertainty bounds.

What are you going to do now?

- Remember that high-dimensional Bayesian inference can be possible.
- Read the paper and look at the code on GitHub.

New pre-print extending the idea

Correlated pseudo-marginal Metropolis-Hastings using quasi-Newton proposals

Johan Dahlin, Adrian Wills and Brett Ninness*

June 26, 2018

Abstract

Pseudo-marginal Metropolis-Hastings (pmMH) is a versatile algorithm for sampling from target distributions which are not easy to evaluate point-wise. However, pmMH requires good proposal distributions to sample efficiently from the target, which can be problematic to construct in practice. This is especially a problem for high-dimensional targets when the standard random-walk proposal is inefficient.

arXiv pre-print: 1806.09780.

Thank you for listening

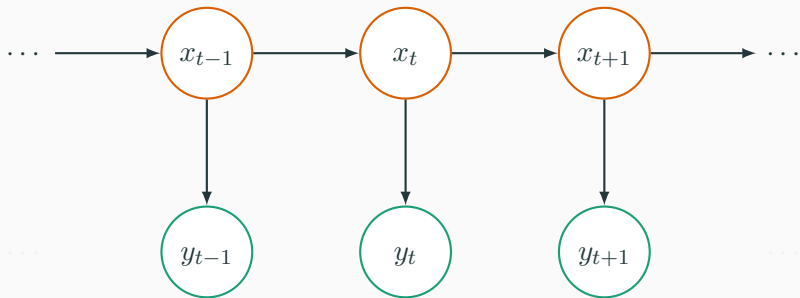
Comments, suggestions and/or questions?

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State-space models

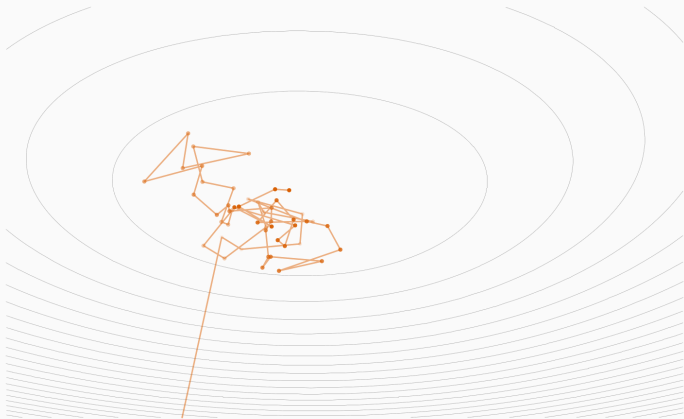


$$x_0 \sim \mu_\theta(x_0) \quad x_{t+1}|x_t \sim f_\theta(x_{t+1}|x_t), \quad y_t|x_t \sim g_\theta(y_t|x_t).$$

Linear Gaussian state-space model:

$$x_{t+1}|x_t \sim \mathcal{N}(x_{t+1}; \mu + \phi(x_t - \mu), \sigma_v^2), \quad y_t|x_t \sim \mathcal{N}(y_t; x_t, \sigma_e^2).$$

Noisy gradient ascent update



$$\theta_k | \theta_{k-1} \sim \mathcal{N}\left(\theta_k; \theta_{k-1} + \frac{1}{2} \Sigma G(\theta_{k-1}), \Sigma\right).$$

Logit function

The logit function is given by

$$\text{logit}(f(x)) = \frac{1}{1 + \exp(-f(x))}$$

and squeezes a real-valued number into the unit interval:

